

Lecture 13: Some Important Continuous Probability Distributions (Part 2)

Kateřina Staňková

Statistics (MAT1003)

May 10, 2012

Outline

- 1 **Flavor of estimation problems ...**
- 2 **Exponential Distribution**
 - Formulation
 - Expectation etc.
 - Application of the Exponential distribution
- 3 **Normal Distribution**
 - Basics
 - Examples
- 4 **Exercises**
- 5 **Monday**

book: Sections 6.2-6.4,6.6

And now ...

- 1 **Flavor of estimation problems ...**
- 2 **Exponential Distribution**
 - Formulation
 - Expectation etc.
 - Application of the Exponential distribution
- 3 **Normal Distribution**
 - Basics
 - Examples
- 4 **Exercises**
- 5 **Monday**

Example 3: Estimation in a uniform distribution $U(0, L)$

Let $X \sim U(0, L)$, with L unknown

Example 3: Estimation in a uniform distribution $U(0, L)$

Let $X \sim U(0, L)$, with L unknown

- We want to estimate L

Example 3: Estimation in a uniform distribution $U(0, L)$

Let $X \sim U(0, L)$, with L unknown

- We want to estimate L
- For that purpose we draw $100 \times$ independently from $U(0, L)$

Example 3: Estimation in a uniform distribution $U(0, L)$

Let $X \sim U(0, L)$, with L unknown

- We want to estimate L
- For that purpose we draw $100 \times$ independently from $U(0, L)$
- Let X_i be the RV corresponding to the i^{th} drawing

Example 3: Estimation in a uniform distribution $U(0, L)$

Let $X \sim U(0, L)$, with L unknown

- We want to estimate L
- For that purpose we draw $100 \times$ independently from $U(0, L)$
- Let X_i be the RV corresponding to the i^{th} drawing
- Then X_1, X_2, \dots, X_{100} are independent and identically distributed (IID) according to $U(0, L)$

Example 3: Estimation in a uniform distribution $U(0, L)$

Let $X \sim U(0, L)$, with L unknown

- We want to estimate L
- For that purpose we draw $100 \times$ independently from $U(0, L)$
- Let X_i be the RV corresponding to the i^{th} drawing
- Then X_1, X_2, \dots, X_{100} are independent and identically distributed (IID) according to $U(0, L)$
- We call X_1, X_2, \dots, X_{100} a random sample

Example 3: Estimation in a uniform distribution $U(0, L)$

Let $X \sim U(0, L)$, with L unknown

- We want to estimate L
- For that purpose we draw $100 \times$ independently from $U(0, L)$
- Let X_i be the RV corresponding to the i^{th} drawing
- Then X_1, X_2, \dots, X_{100} are independent and identically distributed (IID) according to $U(0, L)$
- We call X_1, X_2, \dots, X_{100} a random sample
- After 100 drawings we have 100 realizations, denoted by X_1, X_2, \dots, X_{100}

Example 3: Estimation in a uniform distribution $U(0, L)$ (cont.)

Example 3: Estimation in a uniform distribution $U(0, L)$ (cont.)

- Define $Z = \max\{X_1, \dots, X_{100}\}$

Example 3: Estimation in a uniform distribution $U(0, L)$ (cont.)

- Define $Z = \max\{X_1, \dots, X_{100}\}$
- Then $F(z) = P(Z \leq z) = \prod_{i=1}^{100} P(X_i \leq z) = \left(\frac{z}{L}\right)^{100}, 0 \leq z \leq L$

Example 3: Estimation in a uniform distribution $U(0, L)$ (cont.)

- Define $Z = \max\{X_1, \dots, X_{100}\}$
- Then $F(z) = P(Z \leq z) = \prod_{i=1}^{100} P(X_i \leq z) = \left(\frac{z}{L}\right)^{100}, 0 \leq z \leq L$
- Hence $f(z) = F'(z) = 100 \cdot \left(\frac{z}{L}\right)^{99} \cdot \frac{1}{L}$

Example 3: Estimation in a uniform distribution $U(0, L)$ (cont.)

- Define $Z = \max\{X_1, \dots, X_{100}\}$
- Then $F(z) = P(Z \leq z) = \prod_{i=1}^{100} P(X_i \leq z) = \left(\frac{z}{L}\right)^{100}$, $0 \leq z \leq L$
- Hence $f(z) = F'(z) = 100 \cdot \left(\frac{z}{L}\right)^{99} \cdot \frac{1}{L}$
- Then $E(Z) = \int_0^L z \cdot 100 \cdot \left(\frac{z}{L}\right)^{99} \cdot \frac{1}{L} dz = 100 \int_0^L \left(\frac{z}{L}\right)^{100} dz$
 $= \left[100 \cdot \frac{1}{101} \left(\frac{z}{L}\right)^{101} \cdot L\right]_{z=0}^L = \frac{100}{101} L$

Example 3: Estimation in a uniform distribution $U(0, L)$ (cont.)

- Define $Z = \max\{X_1, \dots, X_{100}\}$
- Then $F(z) = P(Z \leq z) = \prod_{i=1}^{100} P(X_i \leq z) = \left(\frac{z}{L}\right)^{100}$, $0 \leq z \leq L$
- Hence $f(z) = F'(z) = 100 \cdot \left(\frac{z}{L}\right)^{99} \cdot \frac{1}{L}$
- Then $E(Z) = \int_0^L z \cdot 100 \cdot \left(\frac{z}{L}\right)^{99} \cdot \frac{1}{L} dz = 100 \int_0^L \left(\frac{z}{L}\right)^{100} dz$
 $= \left[100 \cdot \frac{1}{101} \left(\frac{z}{L}\right)^{101} \cdot L\right]_{z=0}^L = \frac{100}{101} L$
- As an estimate for L we now define the RV $B = \frac{101}{100} \cdot Z$. We have $E(B) = L$

Example 3: Estimation in a uniform distribution $U(0, L)$ (cont.)

- Define $Z = \max\{X_1, \dots, X_{100}\}$
- Then $F(z) = P(Z \leq z) = \prod_{i=1}^{100} P(X_i \leq z) = \left(\frac{z}{L}\right)^{100}$, $0 \leq z \leq L$
- Hence $f(z) = F'(z) = 100 \cdot \left(\frac{z}{L}\right)^{99} \cdot \frac{1}{L}$
- Then $E(Z) = \int_0^L z \cdot 100 \cdot \left(\frac{z}{L}\right)^{99} \cdot \frac{1}{L} dz = 100 \int_0^L \left(\frac{z}{L}\right)^{100} dz$
 $= \left[100 \cdot \frac{1}{101} \left(\frac{z}{L}\right)^{101} \cdot L\right]_{z=0}^L = \frac{100}{101} L$
- As an estimate for L we now define the RV $B = \frac{101}{100} \cdot Z$. We have $E(B) = L$
- B is called an **unbiased estimator** for L

Estimates

Estimates

- **Point estimate** - solution is a single point

Estimates

- **Point estimate** - solution is a single point
- **Interval estimate** - solution is an interval

Estimates

- **Point estimate** - solution is a single point
- **Interval estimate** - solution is an interval

2 common point estimates

- **The sample mean** - $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$

Estimates

- **Point estimate** - solution is a single point
- **Interval estimate** - solution is an interval

2 common point estimates

- **The sample mean** - $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$
- **The sample variance** - $\bar{S}^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

Estimates

- **Point estimate** - solution is a single point
- **Interval estimate** - solution is an interval

2 common point estimates

- The **sample mean** - $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$
- The **sample variance** - $\bar{S}^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

Observation 1

- $E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \sum_{i=1}^n \mu_X = \frac{1}{n} \cdot n \cdot \mu_X = \mu_X$

Observation 2

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$V(\bar{X}) = V\left(\frac{1}{n} \sum_{i=1}^n X_i\right)$$

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$V(\bar{X}) = V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = V\left(\frac{1}{n} X_1 + \frac{1}{n} X_2 + \dots + \frac{1}{n} X_n\right)$$

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$\begin{aligned}V(\bar{X}) &= V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = V\left(\frac{1}{n} X_1 + \frac{1}{n} X_2 + \dots + \frac{1}{n} X_n\right) \\ &= \frac{1}{n^2} V(X_1) + \frac{1}{n^2} V(X_2) + \dots + \frac{1}{n^2} V(X_n)\end{aligned}$$

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$\begin{aligned}V(\bar{X}) &= V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = V\left(\frac{1}{n} X_1 + \frac{1}{n} X_2 + \dots + \frac{1}{n} X_n\right) \\&= \frac{1}{n^2} V(X_1) + \frac{1}{n^2} V(X_2) + \dots + \frac{1}{n^2} V(X_n) \\&= \frac{1}{n^2} \cdot n \cdot V(X_i)\end{aligned}$$

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$\begin{aligned}V(\bar{X}) &= V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = V\left(\frac{1}{n} X_1 + \frac{1}{n} X_2 + \dots + \frac{1}{n} X_n\right) \\&= \frac{1}{n^2} V(X_1) + \frac{1}{n^2} V(X_2) + \dots + \frac{1}{n^2} V(X_n) \\&= \frac{1}{n^2} \cdot n \cdot V(X_i) = \frac{1}{n} \cdot V(X_i)\end{aligned}$$

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$\begin{aligned}V(\bar{X}) &= V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = V\left(\frac{1}{n} X_1 + \frac{1}{n} X_2 + \dots + \frac{1}{n} X_n\right) \\&= \frac{1}{n^2} V(X_1) + \frac{1}{n^2} V(X_2) + \dots + \frac{1}{n^2} V(X_n) \\&= \frac{1}{n^2} \cdot n \cdot V(X_i) = \frac{1}{n} \cdot V(X_i) = \frac{1}{n} \sigma^2\end{aligned}$$

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$\begin{aligned}V(\bar{X}) &= V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = V\left(\frac{1}{n} X_1 + \frac{1}{n} X_2 + \dots + \frac{1}{n} X_n\right) \\&= \frac{1}{n^2} V(X_1) + \frac{1}{n^2} V(X_2) + \dots + \frac{1}{n^2} V(X_n) \\&= \frac{1}{n^2} \cdot n \cdot V(X_i) = \frac{1}{n} \cdot V(X_i) = \frac{1}{n} \sigma^2\end{aligned}$$

Notice that $\underbrace{V(\bar{X}) = E\{(X - \mu)^2\}}_{\text{variance of sample mean}} \neq \underbrace{\bar{S}^2 \approx E\{(X_i - \bar{X})^2\}}_{\text{sample variance}}$

Observation 2

Let X_1, \dots, X_n be IID with $\mu = E(X_i)$, $\sigma^2 = V(X_i)$. Then

$$\begin{aligned}V(\bar{X}) &= V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = V\left(\frac{1}{n} X_1 + \frac{1}{n} X_2 + \dots + \frac{1}{n} X_n\right) \\&= \frac{1}{n^2} V(X_1) + \frac{1}{n^2} V(X_2) + \dots + \frac{1}{n^2} V(X_n) \\&= \frac{1}{n^2} \cdot n \cdot V(X_i) = \frac{1}{n} \cdot V(X_i) = \frac{1}{n} \sigma^2\end{aligned}$$

Notice that $\underbrace{V(\bar{X}) = E\{(X - \mu)^2\}}_{\text{variance of sample mean}} \neq \underbrace{\bar{S}^2 \approx E\{(X_i - \bar{X})^2\}}_{\text{sample variance}}$

Moreover, Observation 2 is **independent** of the actual distribution of the X_i

Example 4

Example 4

Let X have the following distribution: $P(X = 1) = \bar{p}$, $P(X = 0) = 1 - \bar{p}$ with \bar{p} unknown (0 elsewhere). Estimate \bar{p}

Example 4

Let X have the following distribution: $P(X = 1) = \bar{p}$, $P(X = 0) = 1 - \bar{p}$ with \bar{p} unknown (0 elsewhere). Estimate \bar{p}

- Notice that $\mu_X = E(X) = \bar{p}$, $V(X) = \bar{p} - \bar{p}^2 = \bar{p}(1 - \bar{p})$

Example 4

Let X have the following distribution: $P(X = 1) = \bar{p}$, $P(X = 0) = 1 - \bar{p}$ with \bar{p} unknown (0 elsewhere). Estimate \bar{p}

- Notice that $\mu_X = E(X) = \bar{p}$, $V(X) = \bar{p} - \bar{p}^2 = \bar{p}(1 - \bar{p})$
- We draw X_1, \dots, X_n from this distribution. Then:

Example 4

Let X have the following distribution: $P(X = 1) = \bar{p}$, $P(X = 0) = 1 - \bar{p}$ with \bar{p} unknown (0 elsewhere). Estimate \bar{p}

- Notice that $\mu_X = E(X) = \bar{p}$, $V(X) = \bar{p} - \bar{p}^2 = \bar{p}(1 - \bar{p})$
- We draw X_1, \dots, X_n from this distribution. Then:
 - $E(\bar{X}) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \cdot n \cdot E(X) = \bar{p}$

Example 4

Let X have the following distribution: $P(X = 1) = \bar{p}$, $P(X = 0) = 1 - \bar{p}$ with \bar{p} unknown (0 elsewhere). Estimate \bar{p}

- Notice that $\mu_X = E(X) = \bar{p}$, $V(X) = \bar{p} - \bar{p}^2 = \bar{p}(1 - \bar{p})$
- We draw X_1, \dots, X_n from this distribution. Then:
 - $E(\bar{X}) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \cdot n \cdot E(X) = \bar{p} \Rightarrow \bar{X}$ unbiased estimator for $\mu_X = \bar{p}$

Example 4

Let X have the following distribution: $P(X = 1) = \bar{p}$, $P(X = 0) = 1 - \bar{p}$ with \bar{p} unknown (0 elsewhere). Estimate \bar{p}

- Notice that $\mu_X = E(X) = \bar{p}$, $V(X) = \bar{p} - \bar{p}^2 = \bar{p}(1 - \bar{p})$
- We draw X_1, \dots, X_n from this distribution. Then:
 - $E(\bar{X}) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \cdot n \cdot E(X) = \bar{p} \Rightarrow \bar{X}$ unbiased estimator for $\mu_X = \bar{p}$
 - $V(\bar{X}) = \frac{1}{n^2} \sum_{i=1}^n V(X_i) = \frac{1}{n^2} \cdot n \cdot \bar{p}(1 - \bar{p}) = \frac{1}{n} \cdot \bar{p}(1 - \bar{p})$

Example 4

Let X have the following distribution: $P(X = 1) = \bar{p}$, $P(X = 0) = 1 - \bar{p}$ with \bar{p} unknown (0 elsewhere). Estimate \bar{p}

- Notice that $\mu_X = E(X) = \bar{p}$, $V(X) = \bar{p} - \bar{p}^2 = \bar{p}(1 - \bar{p})$
 - We draw X_1, \dots, X_n from this distribution. Then:
 - $E(\bar{X}) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \cdot n \cdot E(X) = \bar{p} \Rightarrow \bar{X}$ unbiased estimator for $\mu_X = \bar{p}$
 - $V(\bar{X}) = \frac{1}{n^2} \sum_{i=1}^n V(X_i) = \frac{1}{n^2} \cdot n \cdot \bar{p}(1 - \bar{p}) = \frac{1}{n} \cdot \bar{p}(1 - \bar{p})$
- HW: What is the probability distribution of $\sum_{i=1}^n X_i$?

And now ...

- 1 Flavor of estimation problems ...
- 2 Exponential Distribution**
 - Formulation
 - Expectation etc.
 - Application of the Exponential distribution
- 3 Normal Distribution
 - Basics
 - Examples
- 4 Exercises
- 5 Monday

Formulation

What is exponential PDF?

A RV X is exponentially distributed with parameter λ (book: parameter $1/\beta$), $X \sim \text{Exp}(\lambda)$, if

What is exponential PDF?

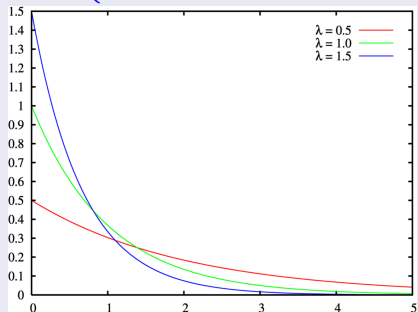
A RV X is exponentially distributed with parameter λ (book: parameter $1/\beta$),
 $X \sim \text{Exp}(\lambda)$, if

$$f(x) = \begin{cases} \lambda \exp(-\lambda x), & x > 0, \\ 0, & \text{elsewhere} \end{cases}$$

What is exponential PDF?

A RV X is exponentially distributed with parameter λ (book: parameter $1/\beta$), $X \sim \text{Exp}(\lambda)$, if

$$f(x) = \begin{cases} \lambda \exp(-\lambda x), & x > 0, \\ 0, & \text{elsewhere} \end{cases}$$



Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$E(X)$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$E(X) = \int_0^{\infty} x f(x) dx$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$E(X) = \int_0^{\infty} x f(x) dx = \int_0^{\infty} x \cdot \lambda \exp(-\lambda x) dx$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$\begin{aligned} E(X) &= \int_0^{\infty} x f(x) dx = \int_0^{\infty} x \cdot \lambda \exp(-\lambda x) dx \\ &\stackrel{bp}{=} [-x \cdot \exp(-\lambda x)]_{x=0}^{\infty} + \int_0^{\infty} \exp(-\lambda x) dx \end{aligned}$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$\begin{aligned} E(X) &= \int_0^{\infty} x f(x) dx = \int_0^{\infty} x \cdot \lambda \exp(-\lambda x) dx \\ &\stackrel{bp}{=} [-x \cdot \exp(-\lambda x)]_{x=0}^{\infty} + \int_0^{\infty} \exp(-\lambda x) dx \\ &= -0 + 0 + \left[-\frac{1}{\lambda} \cdot \exp(-\lambda x) \right]_{x=0}^{\infty} \end{aligned}$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$\begin{aligned} E(X) &= \int_0^{\infty} x f(x) dx = \int_0^{\infty} x \cdot \lambda \exp(-\lambda x) dx \\ &\stackrel{bp}{=} [-x \cdot \exp(-\lambda x)]_{x=0}^{\infty} + \int_0^{\infty} \exp(-\lambda x) dx \\ &= -0 + 0 + \left[-\frac{1}{\lambda} \cdot \exp(-\lambda x) \right]_{x=0}^{\infty} = \frac{1}{\lambda} \end{aligned}$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$E(X) = \int_0^{\infty} x f(x) dx = \int_0^{\infty} x \cdot \lambda \exp(-\lambda x) dx$$

$$\stackrel{bp}{=} [-x \cdot \exp(-\lambda x)]_{x=0}^{\infty} + \int_0^{\infty} \exp(-\lambda x) dx$$

$$= -0 + 0 + \left[-\frac{1}{\lambda} \cdot \exp(-\lambda x) \right]_{x=0}^{\infty} = \frac{1}{\lambda}$$

$$F(x) = 1 - \exp(-\lambda x) \quad \text{for } x > 0 \text{ and } 0 \text{ elsewhere}$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$E(X) = \int_0^{\infty} x f(x) dx = \int_0^{\infty} x \cdot \lambda \exp(-\lambda x) dx$$

$$\stackrel{bp}{=} [-x \cdot \exp(-\lambda x)]_{x=0}^{\infty} + \int_0^{\infty} \exp(-\lambda x) dx$$

$$= -0 + 0 + \left[-\frac{1}{\lambda} \cdot \exp(-\lambda x) \right]_{x=0}^{\infty} = \frac{1}{\lambda}$$

$$F(x) = 1 - \exp(-\lambda x) \quad \text{for } x > 0 \text{ and } 0 \text{ elsewhere}$$

$$E(X^2) = \frac{2}{\lambda^2}$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Computation of $E(X)$, $F(X)$ and $V(X)$ for the Exponential distribution

$$E(X) = \int_0^{\infty} x f(x) dx = \int_0^{\infty} x \cdot \lambda \exp(-\lambda x) dx$$

$$\stackrel{bp}{=} [-x \cdot \exp(-\lambda x)]_{x=0}^{\infty} + \int_0^{\infty} \exp(-\lambda x) dx$$

$$= -0 + 0 + \left[-\frac{1}{\lambda} \cdot \exp(-\lambda x) \right]_{x=0}^{\infty} = \frac{1}{\lambda}$$

$$F(x) = 1 - \exp(-\lambda x) \quad \text{for } x > 0 \text{ and } 0 \text{ elsewhere}$$

$$E(X^2) = \frac{2}{\lambda^2}$$

$$V(X) = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

Book: $\beta = \frac{1}{\lambda}$, but λ -notation is more standard ...

Application of the Exponential distribution

What is it good for?

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
 - (b) there is a gap of at least 30 seconds between 2 successive calls?
- (a) $X = \#$ calls in 3 minutes

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)
 $P(X \geq 25) =$

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)
 $P(X \geq 25) = 1 - P(X \leq 25)$

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)
 $P(X \geq 25) = 1 - P(X \leq 25) = 1 - 0.9317$

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)
 $P(X \geq 25) = 1 - P(X \leq 25) = 1 - 0.9317 = 0.0683$

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)

$$P(X \geq 25) = 1 - P(X \leq 25) = 1 - 0.9317 = 0.0683$$

(b) Y : interarrival time between two calls \Rightarrow

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)

$$P(X \geq 25) = 1 - P(X \leq 25) = 1 - 0.9317 = 0.0683$$

(b) Y : interarrival time between two calls $\Rightarrow Y \sim \text{Exp}(\lambda = 18)$ per 3 minutes,

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)

$$P(X \geq 25) = 1 - P(X \leq 25) = 1 - 0.9317 = 0.0683$$

(b) Y : interarrival time between two calls $\Rightarrow Y \sim \text{Exp}(\lambda = 18)$ per 3 minutes, ($Y \sim \text{Exp}(\lambda = 3)$ per 1/2 minute)

$$P(Y \geq 1/6) = \int_{1/6}^{\infty} 18 \cdot \exp(-18x) dx = e^{-3}$$

Application of the Exponential distribution

What is it good for?

- Imagine X distributed according to **Poisson Process**, i.e., $X \sim \mathcal{P}(\mu)$, i.e., we have on average μ of arrivals per time unit
- Then the time between 2 arrivals, the so-called **interarrival time** is exponentially distributed with parameter $\lambda = \mu$ (book: $\beta = \frac{1}{\mu}$)

Example 5

Suppose on average 6 people call some service number per minute. What is the probability that:

- (a) in the next 3 minutes at least 25 people call?
- (b) there is a gap of at least 30 seconds between 2 successive calls?

(a) $X = \#$ calls in 3 minutes $\Rightarrow X \sim \mathcal{P}(18)$ minutes (time unit: 3 minutes here)

$$P(X \geq 25) = 1 - P(X \leq 25) = 1 - 0.9317 = 0.0683$$

(b) Y : interarrival time between two calls $\Rightarrow Y \sim \text{Exp}(\lambda = 18)$ per 3 minutes, ($Y \sim \text{Exp}(\lambda = 3)$ per 1/2 minute)

$$P(Y \geq 1/6) = \int_{1/6}^{\infty} 18 \cdot \exp(-18x) dx = e^{-3}$$

(Similarly with $\lambda = 3$: $P(Y \geq 1) = e^{-3}$)

And now ...

- 1 Flavor of estimation problems ...
- 2 Exponential Distribution
 - Formulation
 - Expectation etc.
 - Application of the Exponential distribution
- 3 Normal Distribution
 - Basics
 - Examples
- 4 Exercises
- 5 Monday

Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

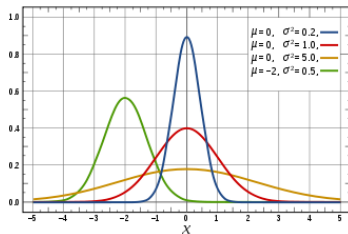
$x \in \mathbb{R}$

Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$



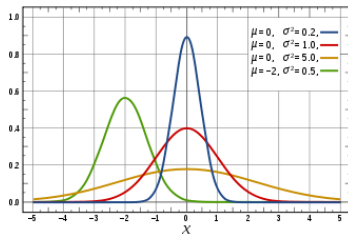
Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$

Notice:



Basics

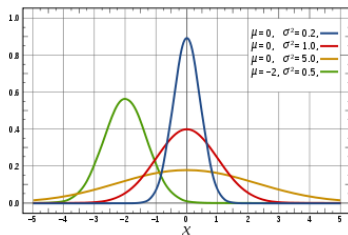
An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$

Notice:

- f is symmetric around $x = \mu$

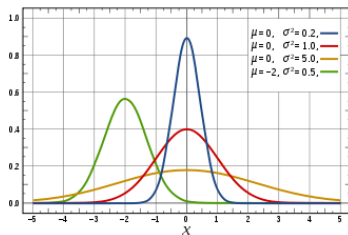


Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$



Notice:

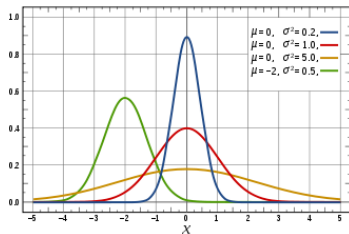
- f is symmetric around $x = \mu$
- f is maximal for $x = \mu$

Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$



Notice:

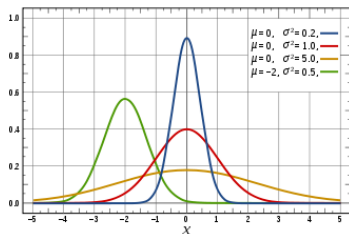
- f is symmetric around $x = \mu$
- f is maximal for $x = \mu$
- One can prove that $E(X) = \mu, V(X) = \sigma^2$

Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$$x \in \mathbb{R}$$



Notice:

- f is symmetric around $x = \mu$
- f is maximal for $x = \mu$
- One can prove that $E(X) = \mu, V(X) = \sigma^2$

Theorem: $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$

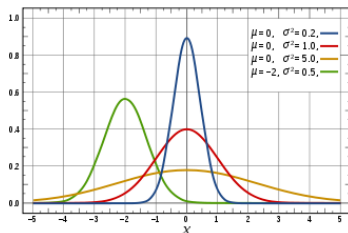
Proof:

Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$



Notice:

- f is symmetric around $x = \mu$
- f is maximal for $x = \mu$
- One can prove that $E(X) = \mu$, $V(X) = \sigma^2$

Theorem: $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$

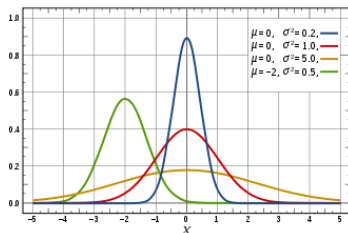
Proof: Let $\tilde{z} = \frac{\tilde{x}-\mu}{\sigma}$. Then $P(X \leq \tilde{x}) = P\left(\frac{X-\mu}{\sigma} \leq \frac{\tilde{x}-\mu}{\sigma}\right) = P(Z \leq \tilde{z})$

Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$



Notice:

- f is symmetric around $x = \mu$
- f is maximal for $x = \mu$
- One can prove that $E(X) = \mu$, $V(X) = \sigma^2$

Theorem: $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$

Proof: Let $\tilde{x} = \frac{x-\mu}{\sigma}$. Then $P(X \leq \tilde{x}) = P\left(\frac{X-\mu}{\sigma} \leq \frac{\tilde{x}-\mu}{\sigma}\right) = P(Z \leq \tilde{z})$

$$P(Z \leq \tilde{z}) = P(X \leq \tilde{x}) = F(\tilde{x}) = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\tilde{x}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right) dx$$

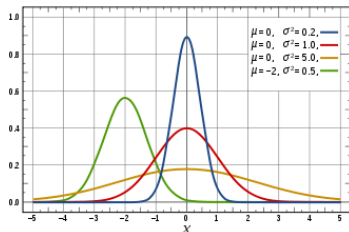
$$\stackrel{z = \frac{x-\mu}{\sigma}, dz = \frac{1}{\sigma} dx}{=} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\tilde{z}} \exp\left(-\frac{1}{2}z^2\right) dz$$

Basics

An RV X has a normal distribution with parameters μ and σ ($X \sim \mathcal{N}(\mu, \sigma)$), if the density is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$$

$x \in \mathbb{R}$



Notice:

- f is symmetric around $x = \mu$
- f is maximal for $x = \mu$
- One can prove that $E(X) = \mu$, $V(X) = \sigma^2$

Theorem: $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$

Proof: Let $\tilde{z} = \frac{\tilde{x}-\mu}{\sigma}$. Then $P(X \leq \tilde{x}) = P\left(\frac{X-\mu}{\sigma} \leq \frac{\tilde{x}-\mu}{\sigma}\right) = P(Z \leq \tilde{z})$

$$P(Z \leq \tilde{z}) = P(X \leq \tilde{x}) = F(\tilde{x}) = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\tilde{x}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right) dx$$

$$z = \frac{x-\mu}{\sigma}, dz = \frac{1}{\sigma} dx$$

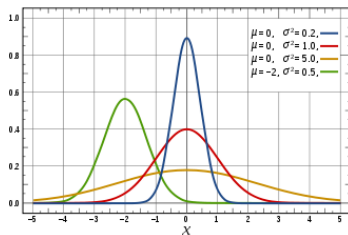
$$= \frac{1}{2\pi} \int_{-\infty}^{\tilde{z}} \exp\left(-\frac{1}{2}z^2\right) dz$$

$$f(\tilde{z}) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}\tilde{z}^2\right), \text{ thus } Z \sim \mathcal{N}(0, 1) \quad \square$$

Basics

If $X \sim \mathcal{N}(\mu, \sigma)$, then

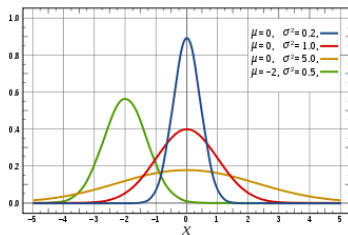
- $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$
- $E(X) = \mu, V(X) = \sigma^2$
- $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$



Basics

If $X \sim \mathcal{N}(\mu, \sigma)$, then

- $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$
- $E(X) = \mu, V(X) = \sigma^2$
- $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$

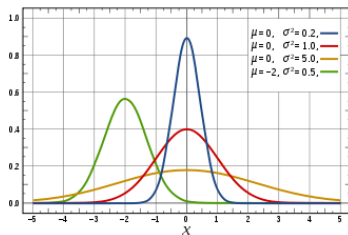


Basics

If $X \sim \mathcal{N}(\mu, \sigma)$, then

- $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$
- $E(X) = \mu, V(X) = \sigma^2$
- $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$

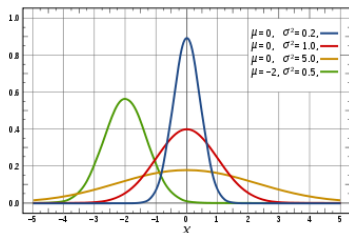
- Table A3: $P(Z \leq z)$ ($Z \sim \mathcal{N}(0, 1)$)



Basics

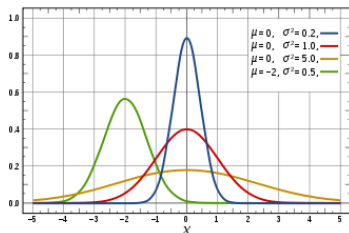
If $X \sim \mathcal{N}(\mu, \sigma)$, then

- $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$
- $E(X) = \mu$, $V(X) = \sigma^2$
- $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$
- Table A3: $P(Z \leq z)$ ($Z \sim \mathcal{N}(0, 1)$)
- If $X_1 \sim \mathcal{N}(\mu_1, \sigma_1)$ and $X_2 \sim \mathcal{N}(\mu_2, \sigma_2)$ and X_1 and X_2 are independent, then: $X_1 + X_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sqrt{\sigma_1^2 + \sigma_2^2})$, $V(X_1 + X_2) = V(X_1) + V(X_2)$



If $X \sim \mathcal{N}(\mu, \sigma)$, then

- $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$
- $E(X) = \mu$, $V(X) = \sigma^2$
- $X \sim \mathcal{N}(\mu, \sigma)$ and $Z = \frac{X-\mu}{\sigma}$, then $Z \sim \mathcal{N}(0, 1)$



- Table A3: $P(Z \leq z)$ ($Z \sim \mathcal{N}(0, 1)$)
- If $X_1 \sim \mathcal{N}(\mu_1, \sigma_1)$ and $X_2 \sim \mathcal{N}(\mu_2, \sigma_2)$ and X_1 and X_2 are independent, then: $X_1 + X_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sqrt{\sigma_1^2 + \sigma_2^2})$, $V(X_1 + X_2) = V(X_1) + V(X_2)$
- If $X_i \sim \mathcal{N}(\mu, \sigma)$, $i = 1, \dots, n$, X_1, \dots, X_n IID. Then
 - $\sum_{i=1}^n X_i \sim \mathcal{N}(n \cdot \mu, \sqrt{n} \cdot \sigma)$
 - $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \sim \mathcal{N}(\mu, \sigma/\sqrt{n})$, $E(\bar{X}) = \mu$

Example 6(a)

The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ g and $\sigma = 5$ g. What is the probability that the net weight of a pack is at least 500 g?

Example 6(a)

The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ g and $\sigma = 5$ g. What is the probability that the net weight of a pack is at least 500 g?

$$P(X \geq 500) = P\left(Z \geq \frac{500 - \mu}{\sigma} = \frac{500 - 505}{5} = -1\right)$$

Example 6(a)

The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ g and $\sigma = 5$ g. What is the probability that the net weight of a pack is at least 500 g?

$$\begin{aligned} P(X \geq 500) &= P\left(Z \geq \frac{500 - \mu}{\sigma} = \frac{500 - 505}{5} = -1\right) \\ &= 1 - P(Z \leq -1) \stackrel{\text{Table A3}}{=} 1 - 0.1587 = 0.8413 \end{aligned}$$

And now ...

- 1 Flavor of estimation problems ...
- 2 **Exponential Distribution**
 - Formulation
 - Expectation etc.
 - Application of the Exponential distribution
- 3 **Normal Distribution**
 - Basics
 - Examples
- 4 **Exercises**
- 5 **Monday**

Computing together:

- book (pp. 164-165): 5.51, 5.59, 5.65
- book (pp. 186-187): 6.5, 6.7, 6.13

And now ...

- 1 Flavor of estimation problems ...
- 2 Exponential Distribution
 - Formulation
 - Expectation etc.
 - Application of the Exponential distribution
- 3 Normal Distribution
 - Basics
 - Examples
- 4 Exercises
- 5 **Monday**

- Finishing up continuous PD's, introducing
 - Erlang distribution
 - Gamma-distribution
 - Chi-squared distribution
- Central limit theorem