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Lecture 13: Some Important Continuous Probability Distributions (Part 2)

Kateřina Staňková

Statistics (MAT1003)

May 14, 2012

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Outline



book: Sections 6.2-6.4,6.6

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And now ...



- Formulation
- Application of Erlang distribution
- Gamma distribution

2 Various Exercises

- Chi-squared distribution
 - Basics
 - Applications
 - Examples

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Formulation

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What is Erlang PDF?

Let $X_1, X_2, ..., X_n$ be IID exponentially distributed RV's with parameter λ . Then the RV $X \stackrel{\text{def}}{=} X_1 + X_2 + ... + X_n$ has an Erlang distribution with parameters nand $\lambda, X \sim \text{Erl}(n, \lambda)$, Formulation

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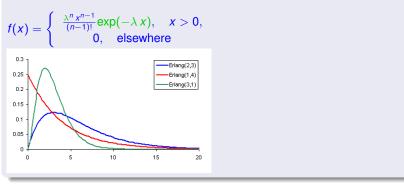
$$f(x) = \begin{cases} \frac{\lambda^n x^{n-1}}{(n-1)!} \exp(-\lambda x), & x > 0, \\ 0, & \text{elsewhere} \end{cases}$$

Formulation

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Application of Erlang distribution

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Application of Erlang distribution

In a Poisson process the sum of *n* interarrival times has an Erlang distribution with parameters *n* and λ

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Example 5(c) (from before)

In a Poisson process the sum of *n* interarrival times has an Erlang distribution with parameters *n* and λ

Example 5(c) (from before)

Suppose on average 6 people call some service number per minute. What is the probability that it takes at least one minute for 3 people to call?

• Z = 3 interarrival times

In a Poisson process the sum of *n* interarrival times has an Erlang distribution with parameters *n* and λ

Example 5(c) (from before)

- Z = 3 interarrival times
- Z ~ Erl(3, 18)

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Example 5(c) (from before)

- Z = 3 interarrival times
- $Z \sim \text{Erl}(3, 18)$ ($\lambda = 18$ for 3-minute time-unit)

$$P(Z \geq \frac{1}{3})$$

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- Z = 3 interarrival times
- $Z \sim \text{Erl}(3, 18)$ ($\lambda = 18$ for 3-minute time-unit)

$$P(Z \ge \frac{1}{3}) = \int_{\frac{1}{3}}^{\infty} \frac{\lambda^n x^{n-1}}{(n-1)!} \exp(-\lambda x) \, \mathrm{d}x$$

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- Z = 3 interarrival times
- *Z* ~ Erl(3, 18) (λ = 18 for 3-minute time-unit)

$$P(Z \ge \frac{1}{3}) = \int_{\frac{1}{3}}^{\infty} \frac{\lambda^n x^{n-1}}{(n-1)!} \exp(-\lambda x) \, \mathrm{d}x = \int_{\frac{1}{3}}^{\infty} \frac{18^3 \cdot x^2}{2!} \exp(-18x) \, \mathrm{d}x$$

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$$= \left[\frac{18^3}{2} \cdot \left(-\frac{1}{18}\right) x^2 \exp(-18 x)\right]_{x=\frac{1}{3}}^{\infty}$$
$$+ \frac{18^3}{2} \int_{\frac{1}{3}}^{\infty} \frac{2}{18} x \exp(-18 x) \, \mathrm{d}x = \left[-162 \, x^2 \exp(-18 x)\right]_{x=\frac{1}{3}}^{\infty}$$

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$$+ \left[324 \cdot \left(-\frac{1}{18}\right) x \exp(-18 x)\right]_{x=\frac{1}{3}}^{\infty} + 18 \int_{\frac{1}{3}}^{\infty} \exp(-18 x) \, dx$$

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$$= \left[18 \exp(-6) + 6 \exp(-6) - \exp(-18 x)\right]_{x=\frac{1}{3}}^{\infty}$$

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Derivation

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is a PDF. An RV with such a PDF has so-called Gamma distribution with parameters α and λ , $X \sim \Gamma(\lambda, \alpha)$ (Plug in n = 1 or $\alpha = 1$ to obtain the exponential distribution)

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$$\begin{aligned} A &= X + Y \sim \mathcal{N}\left(\mu_X + \mu_Y, \sqrt{\sigma_X^2 + \sigma_Y^2}\right) = \mathcal{N}\left(505 + 10, \sqrt{5^2 + 1^2}\right) \\ &= \mathcal{N}\left(515, \sqrt{26}\right) \Rightarrow Z = \frac{A - 515}{\sqrt{26}} \sim \mathcal{N}(0, 1) \\ P(\frac{510 - 515}{\sqrt{26}} \le Z \le \frac{520 - 515}{\sqrt{26}}) \end{aligned}$$

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$$= \mathcal{N}\left(515, \sqrt{26}\right) \Rightarrow Z = \frac{A - 515}{\sqrt{26}} \sim \mathcal{N}(0, 1)$$
$$P(\frac{510 - 515}{\sqrt{26}} \le Z \le \frac{520 - 515}{\sqrt{26}}) = P(Z \le 0.98) - P(Z \le -0.98)$$

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$$= \mathcal{N}\left(515, \sqrt{26}\right) \Rightarrow Z = \frac{A - 515}{\sqrt{26}} \sim \mathcal{N}(0, 1)$$
$$P(\frac{510 - 515}{\sqrt{26}} \le Z \le \frac{520 - 515}{\sqrt{26}}) = P(Z \le 0.98) - P(Z \le -0.98)$$
$$= 0.6730$$

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Example 6 (from last Thursday)

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The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ grams and $\sigma = 5$ grams.

(c) What is the probability that the gross weight of 10 packs is between 5120 and 5170 grams?

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$$P(5120 \le B \le 5170) = P\left(\frac{5120 - 5150}{\sqrt{260}} \le Z \le \frac{5170 - 5150}{\sqrt{260}}\right)$$

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$$P(5120 \le B \le 5170) = P\left(\frac{5120 - 5150}{\sqrt{260}} \le Z \le \frac{5170 - 5150}{\sqrt{260}}\right)$$
$$= P\left(-1.86 \le Z \le 1.24\right)$$

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The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ grams and $\sigma = 5$ grams.

(c) What is the probability that the gross weight of 10 packs is between 5120 and 5170 grams? *B* : Gross weight of 10 packs, $B = \sum_{i=1}^{10} (X_i + Y_i), \text{ with } X_i \sim \mathcal{N}(505, 5), Y_i \sim \mathcal{N}(10, 1)$ $B \sim \mathcal{N}(10 \, \mu_X + 10 \, \mu_Y, \sqrt{10 \, \sigma_X^2 + 10 \, \sigma_Y^2}) = \mathcal{N}(5150, \sqrt{260})$ $\Rightarrow Z = \frac{B - 5150}{\sqrt{260}} \sim \mathcal{N}(0, 1)$

$$P(5120 \le B \le 5170) = P\left(\frac{5120 - 5150}{\sqrt{260}} \le Z \le \frac{5170 - 5150}{\sqrt{260}}\right)$$
$$= P\left(-1.86 \le Z \le 1.24\right) = P\left(Z \le 1.24\right) - P\left(Z \le -1.86\right)$$

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$$= P\left(-1.86 \le Z \le 1.24\right) = P\left(Z \le 1.24\right) - P\left(Z \le -1.86\right)$$
$$= 0.8925 - 0.0314 = 0.8611$$

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Example 6 (from last Thursday)

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Example 6 (from last Thursday)

The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ grams and $\sigma = 5$ grams.

(d) 80 % of the gross weight of the packs is heavier than a particular weight. What is that weight?

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(d) 80 % of the gross weight of the packs is heavier than a particular weight. What is that weight? $A \sim \mathcal{N}(515, \sqrt{26})$

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 P(A ≥ a) = 0.8. What is a?

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But $P(Z \le z) = 0.2$ for z = -0.84

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In general ...

The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ grams and $\sigma = 5$ grams.

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In general ...

• $X \sim \mathcal{N}(\mu, \sigma)$ and $P(X \leq x) = p, P(Z \leq z = \frac{x-\mu}{\sigma}) = p$

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In general ...

• $X \sim \mathcal{N}(\mu, \sigma)$ and $P(X \leq x) = p, P(Z \leq z = \frac{x-\mu}{\sigma}) = p$

• Look up z in table A3 and then calculate $x = \sigma x + \mu$

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Example 6 (from last Thursday)

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Example 6 (from last Thursday)

The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ grams and $\sigma = 5$ grams.

(e) What is the probability that the difference of net weights of two packs is lower or equal than two grams?

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The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ grams and $\sigma = 5$ grams.

Hence
$$X_1 - X_2 \sim \mathcal{N}(0, \sqrt{\sigma_{X_1}^2 + \sigma_{X_2}^2}) = \mathcal{N}\left(0, \sqrt{25 + 25}\right) = \mathcal{N}\left(0, \sqrt{50}\right)$$

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 $P(-2 \le X_1 - X_2 \le 2) = P(X_1 - X_2 \le 2) - P(X_1 - X_2 \le -2)$

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$$\begin{aligned} & P\left(-2 \le X_1 - X_2 \le 2\right) = P(X_1 - X_2 \le 2) - P(X_1 - X_2 \le -2) \\ & P\left(Z \le \frac{2-0}{\sqrt{50}}\right) - P\left(Z \le \frac{-2-0}{\sqrt{50}}\right) \end{aligned}$$

The net weight of a pack of coffee (500 grams) is a normally distributed RV with parameters $\mu = 505$ grams and $\sigma = 5$ grams.

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$$P(-2 \le X_1 - X_2 \le 2) = P(X_1 - X_2 \le 2) - P(X_1 - X_2 \le -2)$$

$$P\left(Z \le \frac{2-0}{\sqrt{50}}\right) - P\left(Z \le \frac{-2-0}{\sqrt{50}}\right)$$

$$P(Z \le 0.28) - P(Z \le -0.28) = 0.6103 - 0.3897 = 0.2206$$

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Observation

Observation

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Observation

•
$$F(x) = P(X \le x) = P(Z^2 \le x) = P(|z| \le \sqrt{x}) = P(-\sqrt{x} \le Z \le \sqrt{x}) = G(\sqrt{x}) - G(-\sqrt{x})$$

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Observation

Let $Z \sim \mathcal{N}(0, 1)$. For cumulative density of the RV $X = Z^2$ we have:

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$$f(x) = F'(x) = \frac{1}{2}x^{-\frac{1}{2}} \cdot G'(\sqrt{x}) - (-\frac{1}{2})x^{-\frac{1}{2}} \cdot G'(-\sqrt{x})$$

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 $= \frac{1}{2}x^{-\frac{1}{2}} \cdot \frac{1}{\sqrt{2\pi}} \exp(-\frac{x}{2}) + \frac{1}{2}x^{-\frac{1}{2}} \cdot \frac{1}{\sqrt{2\pi}} \exp(-\frac{x}{2})$

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Observation

- $F(x) = P(X \le x) = P(Z^2 \le x) = P(|z| \le \sqrt{x}) = P(-\sqrt{x} \le Z \le \sqrt{x}) = G(\sqrt{x}) G(-\sqrt{x})$ where *G* is the cumulative density of *Z*. So • $f(x) = F'(x) = \frac{1}{2}x^{-\frac{1}{2}} \cdot G'(\sqrt{x}) - (-\frac{1}{2})x^{-\frac{1}{2}} \cdot G'(-\sqrt{x})$ $= \frac{1}{2}x^{-\frac{1}{2}} \cdot \frac{1}{\sqrt{2\pi}} \exp(-\frac{x}{2}) + \frac{1}{2}x^{-\frac{1}{2}} \cdot \frac{1}{\sqrt{2\pi}} \exp(-\frac{x}{2}) = \frac{x^{-\frac{1}{2}} \exp(-\frac{x}{2})}{4\sqrt{\pi}}$ • Apparently: $\Gamma(\frac{1}{2}) = \sqrt{\pi}$ and $X = Z^2$ has a Gamma-distribution with
- Apparently: $\Gamma(\frac{1}{2}) = \sqrt{\pi}$ and $X = Z^2$ has a Gamma-distribution with parameters $\lambda = \frac{1}{2}$ ($\beta = 2$) and $\alpha = \frac{1}{2}$

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And now ...

Erlang distribution

- Formulation
- Application of Erlang distribution
- Gamma distribution

2 Various Exercises

- Chi-squared distribution
 - Basics
 - Applications
 - Examples

Basics

Formulation

Basics

Formulation

If
$$X_1, X_2, ..., X_n$$
 are IID, $X_i \sim \mathcal{N}(0, 1)$, then $\chi^2 \stackrel{\text{def}}{=} X_1^2 + X_2^2 + ... + X_n^2$

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Basics

Formulation

If X_1, X_2, \ldots, X_n are IID, $X_i \sim \mathcal{N}(0, 1)$, then $\chi^2 \stackrel{\text{def}}{=} X_1^2 + X_2^2 + \ldots + X_n^2$

 has a Gamma-distribution with parameters α = n/2 and λ = 1/2 (β = 2)

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Basics

Formulation

If X_1, X_2, \ldots, X_n are IID, $X_i \sim \mathcal{N}(0, 1)$, then $\underline{\chi}^2 \stackrel{\text{def}}{=} X_1^2 + X_2^2 + \ldots + X_n^2$

- has a Gamma-distribution with parameters $\alpha = n/2$ and $\lambda = 1/2$ ($\beta = 2$)
- This distribution is also called a Chi-squared distribution with n degrees of freedom

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Basics

Formulation

If X_1, X_2, \ldots, X_n are IID, $X_i \sim \mathcal{N}(0, 1)$, then $\underline{\chi}^2 \stackrel{\text{def}}{=} X_1^2 + X_2^2 + \ldots + X_n^2$

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Notation $\underline{\chi}^2 \sim \chi_n^2$

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Basics

Formulation

If X_1, X_2, \ldots, X_n are IID, $X_i \sim \mathcal{N}(0, 1)$, then $\underline{\chi}^2 \stackrel{\text{def}}{=} X_1^2 + X_2^2 + \ldots + X_n^2$

• has a Gamma-distribution with parameters $\alpha = n/2$ and $\lambda = 1/2$ ($\beta = 2$)

This distribution is also called a Chi-squared distribution with n degrees of freedom

Notation $\underline{\chi}^2 \sim \chi_n^2$ Book: Table A5: X s.t. $P(\underline{\chi}^2 \ge x) = \alpha$ for several values of v (# degrees of freedom) and α

Applications

Estimation of σ

Applications

Estimation of σ

Let $X_1, X_2, \ldots, X_n \sim \mathcal{N}(\mu, \sigma)$ IDD with known μ and unknown σ



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Applications

Estimation of σ

Let $X_1, X_2, \ldots, X_n \sim \mathcal{N}(\mu, \sigma)$ IDD with known μ and unknown σ Then $Z_i = \frac{X_i - \mu}{\sigma} \sim \mathcal{N}(0, 1)$

Applications

Estimation of σ

Let $X_1, X_2, \ldots, X_n \sim \mathcal{N}(\mu, \sigma)$ IDD with known μ and unknown σ Then $Z_i = \frac{X_i - \mu}{\sigma} \sim \mathcal{N}(0, 1)$ $\Rightarrow Z_i^2 = \frac{(X_i - \mu)^2}{\sigma^2} \sim \chi_i^2$



Applications

Estimation of σ

Let $X_1, X_2, ..., X_n \sim \mathcal{N}(\mu, \sigma)$ IDD with known μ and unknown σ Then $Z_i = \frac{X_i - \mu}{\sigma} \sim \mathcal{N}(0, 1)$ $\Rightarrow Z_i^2 = \frac{(X_i - \mu)^2}{\sigma^2} \sim \chi_i^2$ $\Rightarrow \underline{\chi}^2 \stackrel{\text{def}}{=} \sum_{i=1}^n Z_i^2 = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu)^2 \sim \chi_n^2$

Example 7



Example 7

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• So (fill the realization for χ^2 :) $P(0.831 \le \frac{24.83}{\sigma^2} \le 12.832) = 0.975 - 0.025 = 0.95 \Leftrightarrow P(\sigma^2 \in [1.935, 29.88]) = 0.95$

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Example 7 (b)

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Construct a 95% confidence interval for σ :

Example 7 (b)

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- Table A5: $P(\chi^2 \ge 11.143) = 0.025$ and $P(\chi^2 \ge 0.484) = 0.975$

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95% CI for σ^2 : [24.732/11.143, 24.732/0.484] = [0.220, 51.0999]

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95% CI for σ^2 : [24.732/11.143, 24.732/0.484] = [0.220, 51.0999] 95% CI for σ : [1.490, 7.148] Important!: These calculations use the fact that X_i s are normally distributed. If it is not the case, we cannot use χ^2 -distributions!